



An integrated algorithm for the optimal design of stated choice experiments with partial profiles



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ABSTRACT

Stated choice experiments are conducted to identify the attributes that drive people's preferences when choosing between competing options. They are widely used in transportation in order to support the decision making of companies and governmental authorities. A large number of attributes might increase the complexity of the choice task in a choice experiment, and have a detrimental effect on the quality of the results obtained. In order to reduce the cognitive effort required by the experiment, researchers may resort to experimental designs where the levels of some attributes are held constant within a choice situation. These designs are called partial profile designs. In this paper, we propose an integrated algorithm for the generation of D-optimal designs for stated choice experiments with partial profiles. This algorithm optimizes the set of constant attributes and the levels of the varying attributes simultaneously. An extensive computational experiment shows that the designs produced by the integrated algorithm outperform those obtained by existing algorithms, and match the optimal designs that have been analytically derived for a number of benchmark instances. Additionally, we evaluate the performance of the algorithm under varying experimental conditions and study the structure of the designs generated. We also revisit two stated choice experiments in transportation, and describe how the integrated algorithm could help to improve their designs.

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1. Introduction

Stated choice (SC) experiments are widely used to study how people make choices and to identify the elements that drive people's preferences. They are performed in a wide range of applied fields in order to evaluate the trade-offs that people make when choosing between competing options. In a SC experiment, the options under study are characterized by a set of *attributes*, and each *alternative* or *profile* is described as a combination of attribute levels. Each respondent is presented with several groups of alternatives, called *choice situations*, and asked to choose the alternative of his/her preference in each situation. The main goal of a SC experiment is to estimate the importance of each attribute level from the repeated choices made. By doing so, it is possible to make predictions concerning a population's choice behaviour.

SC experiments are widely used in transportation in order to collect data for the study of travelling behaviour (see, for example, [Tilahun et al. \(2007\)](#), [Hensher and Rose \(2007\)](#) and [Hensher \(2008\)](#)) and for assisting governmental authorities

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in their policy-making activities (see, for example, Wang et al. (2002) and Saleh and Farrell (2005)). Among other things, SC experiments are carried out to study preferences for both air and ground transportation. For example, Ahern and Tapley (2008) studied the preferences of passengers for interurban routes in Ireland. They focused on identifying opportunities for improvement and contrasting the services offered by trains and buses. The SC experiment considered five attributes: cost, journey length, frequency, reliability of the service and the presence of toilet facilities. Bliemer and Rose (2011) studied the preferences of airline customers. Their work was motivated by the fact that flying into main airports is becoming more expensive, both for travellers and airlines. For that reason, there has been an increase in the use of regional airports surrounding major cities. Their main goal was to examine the trade-offs that travellers make between ticket price and overall travel time. The SC experiment involved six attributes: airline, ticket price, departure time, transfer time, egress price and egress time. Bliemer and Rose (2011) also summarized the literature using SC experiments published in the most important transportation journals during the period from January 2000 to August 2009. A total of 64 research papers are listed, involving 61 unique experimental designs. A general overview of the traditional design generation techniques, along with some algorithmic approaches, is given by Rose and Bliemer (2009).

The importance of each attribute is usually quantified using a discrete choice model that is built on the assumption that humans attempt to maximize the total utility when making a choice. In this model, the total utility is expressed as a function of the utilities associated with each attribute. Discrete choice models implicitly assume that humans are willing to make compensatory decisions. In other words, that the negative utility due to undesired levels for some attributes can be compensated for by the positive (or less negative) utility due to preferred levels for some others. The use of non-compensatory decision rules (where, for example, respondents focus on one attribute and ignore the others) is in conflict with this assumption and renders the model's predictions questionable. One main reason why respondents may resort to non-compensatory strategies is the complexity of the decision task. Large numbers of attributes increase the required cognitive effort and might lead the respondents to use simpler decision strategies (Caussade et al., 2005; Hensher, 2006a,b). In order to reduce the complexity of the comparison and to prevent respondents from using simpler strategies, it is possible to hold the levels of some attributes constant in every choice situation. This results in profiles or alternatives that direct the attention only to the subset of attributes the levels of which are allowed to vary. These profiles or alternatives are called *partial profiles* (Chrzan, 2010). The number of varying attributes in a partial profile is called the *profile strength* (Großmann et al., 2006).

Several SC experiments in transportation involve large numbers of attributes. For example, an experiment carried out by Anderson et al. (2006) for the study of tourist parking preferences involved seven attributes. Loo et al. (2006) carried out an experiment to study the preferences of the public light bus industry regarding the introduction of alternative fuel vehicles in Hong Kong. This experiment involved eight attributes related to the vehicle characteristics and seven attributes related to the government support. In other studies, the researchers have reduced or limited the number of attributes considered in their experiments in order to avoid an overwhelming complexity (see, for example, Hunt and Abraham (2007) and Sener et al. (2009)). The study of commute mode choice carried out by Bhat and Sardesai (2006) is an extreme example. During the pilot phase, respondents indicated that the experiment was too burdensome and recommended a reduction in the number of attributes and the number of choice situations. This suggestion encouraged the researchers to consider five attributes instead of the initial eight. More recently, SC experiments with partial profiles have been used in order to avoid the use of non-compensatory decision strategies due to large numbers of attributes. For example, Kupfer et al. (2016) studied the effect of six attributes on the airport choice of air freight service providers in Europe. In order to keep the choice tasks manageable for the respondents, they used partial profiles showing only four of the six attributes in each choice situation. Another example in which a partial profile design has proven useful, is the SC experiment used by Verhetsel et al. (2015). They quantified the impact of the different dimensions of accessibility on the location decision process of logistics companies. This study also involved six attributes, four of which were shown in each choice situation.

In order to design a traditional SC experiment, it is necessary to determine the attribute levels of each alternative. Designing a SC experiments with partial profiles, however, requires making an extra decision: for each choice situation in the design, it is necessary to select the set of attributes the levels of which are held constant. This additional requirement makes generating an optimal design for this kind of SC experiments considerably more challenging. Nearly all the SC experiments with partial profiles that are reported in the literature are designed to estimate a multinomial logit (MNL) model. Although there exist more complex discrete choice models (like, for example, the panel mixed logit model (Train, 2009)) that allow for a more elaborate analysis, the traditional MNL offers two major advantages. First, the computational effort required to evaluate the quality of a design is substantially smaller than that demanded by more complex models (Sándor and Wedel, 2002; Yu et al., 2009). Second, optimal designs generated for the MNL model perform relatively well when it comes to estimating panel mixed logit models (Bliemer and Rose, 2010).

The quality of a design to estimate a MNL model depends on the utilities associated with each attribute level. The exact values of these parameters, however, are not known in advance. Optimal designs can be constructed considering only one specific set of prior parameter values; these designs are referred to as *locally optimal designs* (Huber and Zwerina, 1996). Moreover, from a design perspective, the MNL model can be treated as a regular linear model by assuming that all the parameter values are equal to zero. This assumption implies that the respondents have no preference for any of the attribute levels and all utilities equal zero. Hence, the respondents are equally likely to choose any of the alternatives. The designs produced under this assumption are called *utility-neutral (UN) designs*. A sound alternative to the locally optimal designs

is the Bayesian approach, which accounts for the uncertainty about the true values of the parameters (Sándor and Wedel, 2001). This approach evaluates the quality of a design assuming a prior distribution for the parameter values. It has been shown that designs generated using a Bayesian approach generally outperform locally optimal and UN designs in terms of the quality of the parameter estimates (Sándor and Wedel, 2001; Kessels et al., 2011a,c, 2015). For this reason, Bayesian designs now constitute the state of the art for the generation of optimal designs for SC experiments (Bliemer and Rose, 2010).

A substantial part of the literature related to the generation of SC experiments with partial profiles deals with UN designs: Graßhoff et al. (2004) generate optimal UN designs using orthogonal arrays for SC experiments where all attributes have the same number of levels and where there are two alternatives per choice situation. Großmann et al. (2009) analytically derive UN designs for experiments with two alternatives per choice situation and attributes with different numbers of levels. They focus on experiments with two groups of attributes where the number of attribute levels is fixed for every group. More recently, Großmann et al. (2014) extend their previous work to experiments with three groups of attributes.

The set of optimal designs that can be generated using the approach of Großmann et al. (2009) and Großmann et al. (2014) is restricted. Only designs for experiments with two alternatives per choice situation, specific combinations of attribute levels, and specific numbers of choice situations and constant attributes, can be generated. In response to these limitations, Kessels et al. (2011b) propose a flexible two-stage algorithm for the generation of Bayesian optimal designs. The algorithm is capable of generating designs with any number of choice situations, alternatives per choice situation and constant attributes, and with attributes with any numbers of levels. During its first stage, the algorithm determines the set of constant attributes for each choice situation by generating a structure similar to that of a balanced incomplete block design. By doing so, the algorithm attempts to balance the number of times each attribute is held constant. In its second stage, the algorithm determines the levels of the varying attributes by applying a modified version of the coordinate-exchange algorithm (Meyer and Nachtsheim, 1995).

We believe that the approach of Kessels et al. (2011b) has three major weaknesses:

1. As pointed out by Großmann (2013), the two-stage algorithm is unable to match the quality of the optimal UN designs in the benchmark set of Großmann et al. (2009). This is mainly because the selection of constant attributes of the designs generated by the algorithm is different from that of the optimal designs analytically derived.
2. The choice of the constant attributes and that of the levels of the varying attributes are made independently. More importantly, the choice of the constant attributes never changes during the execution of the algorithm. It is well known from the operations research literature that sequential algorithms (those that make each decision required for solving a problem in a different stage) are prone to generate suboptimal solutions. This is because the optimization performed in each stage is severely affected by the decisions made in previous stages (Salhi and Rand, 1989). Several empirical studies confirm that integrated methods outperform sequential methods for the solution of a wide range of combinatorial optimization problems (Nagy and Salhi, 2007; Basten et al., 2012; Schittekat et al., 2013; Vidal et al., 2013).
3. When generating Bayesian optimal designs, the choice of the constant attributes is made independently from the prior information concerning the model parameters. For example, consider the case in which a SC experiment has two attributes with the same number of levels, but the prior variance of the parameters corresponding to the first attribute is much larger than the prior variance of the parameters corresponding to the second attribute. Due to the difference in uncertainty about the attributes' parameters, it is reasonable to expect the attributes to be held constant a different number of times. The two-stage algorithm ignores the prior distribution of parameter values when generating a Bayesian design. Therefore, designs accounting for differences in prior parameter variances in the selection of the constant attributes are not considered by that algorithm.

Kessels et al. (2015) improve the first phase of the two-stage algorithm by incorporating two new strategies for selecting the constant attributes. Instead of paying an equal amount of attention to each attribute, these two strategies pay an equal amount of attention to each attribute *level*. This results in statistically more efficient designs that vary the attributes with large numbers of levels more often than attributes with fewer levels. Nevertheless, the designs generated by the two-stage algorithm are still not able to match the quality of those analytically derived by Großmann et al. (2009).

In this paper, we propose an *integrated algorithm* for the generation of optimal designs for SC experiments with partial profiles. This algorithm overcomes the weaknesses of the algorithms of Kessels et al. (2011b, 2015). To this end, it optimizes the set of constant attributes and the levels of the varying attributes simultaneously. The integrated algorithm generates, in a short computing time, designs that outperform those produced by the algorithm of Kessels et al. (2015). However, for some benchmark instances, the integrated algorithm just misses the optimal design and produces a design that is slightly less efficient. To overcome this minor limitation, we also describe a more extensive algorithm that generates the optimal designs for nearly all cases. Furthermore, we discuss the results of an extensive computational experiment that was carried out with two main objectives: first, to assess the performance of the algorithms under different experimental conditions; and second, to analyse the effects that these conditions have on the structure of the designs generated.

This paper is organized as follows. Section 2 reviews the discrete choice model and the optimality criteria considered to generate SC experiments with partial profiles. The original integrated algorithm is described in Section 3 and a more extensive version is presented in Section 4. Section 5 explains the implementation details considered in order to speed up the execution of the algorithms. The computational experiments and the results obtained are discussed in Section 6. Section 7 illustrates the use of SC experiments with partial profiles in transportation. It revisits two different experiments in

the transportation literature, and explains how their designs can be improved by using the integrated algorithm. The final conclusions and some recommendations for future research are presented in Section 8.

2. The multinomial logit model and optimality criteria

In this paper, we limit our focus to the MNL model involving main effects because of the computational and practical advantages it offers, and because the designs that have been analytically derived for SC experiments with partial profiles are optimal for this specific model only. Therefore, considering a different model would not allow for a fair comparison between the algorithms and the theoretical results available. In Section 7, however, we briefly discuss the use of more complex discrete choice models and the implications this entails.

The MNL model defines the utility $U_{s,j}$ that a respondent gives to an alternative $j, j = 1, \dots, J$, in a choice situation $s, s = 1, \dots, S$, as the sum of a systematic and a random component:

$$U_{s,j} = \mathbf{f}'(\mathbf{x}_{s,j})\boldsymbol{\beta} + \varepsilon_{s,j}. \tag{1}$$

The first term in this sum corresponds to the systematic component associated with the levels of the F attributes in the SC experiment: $\mathbf{x}_{s,j}$ is the $F \times 1$ vector of attribute levels for alternative j in choice situation s , $\mathbf{f}(\mathbf{x}_{s,j})$ is the $K \times 1$ vector containing the attribute levels coded into K variables, and $\boldsymbol{\beta}$ is the $K \times 1$ vector of parameter values for the main effects or marginal utilities of the attribute levels. The vector $\boldsymbol{\beta}$ is the same for every respondent, which implies that respondents' preferences are assumed to be homogeneous. The second term in Eq. (1) corresponds to the utility's random component: $\varepsilon_{s,j}$ is an error term which is assumed to be independent and identically Gumbel distributed. The MNL probability that a respondent chooses alternative j in choice situation s is then

$$p_{s,j} = \frac{\exp(\mathbf{f}'(\mathbf{x}_{s,j})\boldsymbol{\beta})}{\sum_{t=1}^J (\exp(\mathbf{f}'(\mathbf{x}_{s,t})\boldsymbol{\beta}))}, \tag{2}$$

where $\boldsymbol{\beta}$ is estimated using maximum likelihood.

The generation of an optimal design \mathbf{X} for a SC experiment involves the generation of S submatrices $\mathbf{X}_s = [\mathbf{x}_{s,1}, \dots, \mathbf{x}_{s,J}]'$ and depends on the Fisher information matrix

$$\mathbf{M}(\mathbf{X}, \boldsymbol{\beta}) = \sum_{s=1}^S \mathbf{f}'(\mathbf{X}_s)(\mathbf{P}_s - \mathbf{p}_s\mathbf{p}_s')\mathbf{f}(\mathbf{X}_s), \tag{3}$$

where $\mathbf{f}(\mathbf{X}_s) = [\mathbf{f}(\mathbf{x}_{s,1}), \dots, \mathbf{f}(\mathbf{x}_{s,J})]'$, $\mathbf{p}_s = [p_{s,1}, \dots, p_{s,J}]'$ and $\mathbf{P}_s = \text{diag}[p_{s,1}, \dots, p_{s,J}]$. The information matrix summarizes how much information a design for a SC experiment contains concerning the respondents' preferences. By making an appropriate choice for the design, the information content of the SC experiment can be maximized. The most common approach to achieve this is to seek the design that maximizes the logarithm of the determinant of the information matrix. This approach is called the \mathcal{D} -optimal design approach. The problem with this approach is that the information matrix depends on the parameter values $\boldsymbol{\beta}$ through the probabilities $p_{s,j}$, and these values are unknown prior to the SC experiment. In order to deal with this dependency, it is possible to consider a single prior guess $\boldsymbol{\beta}_p$ of $\boldsymbol{\beta}$ and generate a so-called locally optimal design. These designs are called \mathcal{D}_p -optimal and maximize the logarithm of the determinant of the information matrix, $\mathcal{D}_p(\mathbf{X}) = \log |\mathbf{M}(\mathbf{X}, \boldsymbol{\beta}_p)|$. This is equivalent to minimizing the \mathcal{D}_p -error of the design, defined as $|\mathbf{M}^{-1}(\mathbf{X}, \boldsymbol{\beta}_p)|^{(1/K)}$.

Utility-neutral (UN) designs are special cases of locally optimal designs where $\boldsymbol{\beta}_p = \mathbf{0}_K$, with $\mathbf{0}_K$ a $K \times 1$ vector of zeroes. Using this prior vector is equivalent to assuming that the respondents have no preference for any of the alternatives. This assumption causes the probabilities in Eq. (2) to be equal to $1/J$ for all the alternatives in a choice situation. The information matrix then reduces to

$$\mathbf{M}(\mathbf{X}) = J^{-1} \sum_{s=1}^S (\mathbf{f}'(\mathbf{X}_s)\mathbf{f}(\mathbf{X}_s) - J^{-1}(\mathbf{f}'(\mathbf{X}_s)\mathbf{1}_J)(\mathbf{1}_J'\mathbf{f}(\mathbf{X}_s))), \tag{4}$$

where $\mathbf{1}_J$ is a $J \times 1$ vector of ones. UN designs that maximize the logarithm of the determinant of the simplified information matrix in Eq. (4), $\mathcal{D}_0(\mathbf{X}) = \log |\mathbf{M}(\mathbf{X}, \mathbf{0}_K)|$, are referred to as \mathcal{D}_0 -optimal designs. These designs also minimize their \mathcal{D}_0 -error, defined as $|\mathbf{M}^{-1}(\mathbf{X}, \mathbf{0}_K)|^{(1/K)}$.

Locally optimal designs might provide imprecise estimates of $\boldsymbol{\beta}$ if the prior guess $\boldsymbol{\beta}_p$ is not close to the true value of the parameter vector $\boldsymbol{\beta}$. This weakness is overcome by the Bayesian design approach, which accounts for the prior uncertainty concerning the parameter values, and generally results in more precise estimates. The Bayesian design approach averages the value of the \mathcal{D} -optimality criterion over a prior distribution of likely parameter values, $\pi(\boldsymbol{\beta})$. It is customary to consider as a prior distribution $\pi(\boldsymbol{\beta})$ the multivariate normal distribution $\mathcal{N}(\boldsymbol{\beta}|\boldsymbol{\beta}_0, \boldsymbol{\Sigma}_0)$ with prior mean $\boldsymbol{\beta}_0$ and prior variance-covariance matrix $\boldsymbol{\Sigma}_0$. A Bayesian \mathcal{D} -optimal design maximizes the logarithm of the determinant of the information matrix averaged over the K -dimensional space \mathcal{R}_K defined by the prior distribution:

$$\mathcal{D}_B(\mathbf{X}) = \int_{\mathcal{R}_K} \log |\mathbf{M}(\mathbf{X}, \boldsymbol{\beta})| \pi(\boldsymbol{\beta}) d\boldsymbol{\beta}. \tag{5}$$

In order to compare two different designs with the same number of choice situations, we use the relative \mathcal{D} -efficiency as a quality measure. The \mathcal{D} -efficiency of a design \mathbf{X} relative to a design \mathbf{X}^* is defined following the suggestions of [Holling and Schwabe \(2011\)](#) as

$$\mathcal{D}\text{-eff}(\mathbf{X}, \mathbf{X}^*) = \exp\left(\frac{\mathcal{D}(\mathbf{X}) - \mathcal{D}(\mathbf{X}^*)}{K}\right), \quad (6)$$

where \mathcal{D} corresponds to the \mathcal{D}_0 -optimality criterion when comparing UN designs, to the \mathcal{D}_p -optimality criterion when comparing locally optimal designs, and to the \mathcal{D}_B -optimality criterion when comparing Bayesian designs. For UN and locally optimal designs, this expression is equivalent to the more common form of the relative efficiency defined as

$$\mathcal{D}\text{-eff}(\mathbf{X}, \mathbf{X}^*) = \left(\frac{|\mathbf{M}(\mathbf{X})|}{|\mathbf{M}(\mathbf{X}^*)|}\right)^{1/K}. \quad (7)$$

3. Integrated algorithm

In this section, we describe the integrated algorithm for the optimal design of SC experiments with partial profiles. The algorithm requires an initial design as an input parameter. This initial design is constructed by randomly choosing the attributes to be held constant for each choice situation, as well as the levels of the varying attributes for each alternative of the choice situation. The initial design is improved by modifying one choice situation at a time. In general terms, the algorithm improves the quality of the choice situations by means of two procedures. The first procedure aims at finding the best set of attributes to be held constant in the choice situation. The second procedure aims at finding the best levels of the varying attributes. These two procedures are executed sequentially in order to find the configuration of the choice situation that produces the best value of the optimality criterion. The iterative process stops when no further improvement can be obtained for any of the design's choice situations.

[Algorithm 1](#) shows the general layout of the integrated algorithm. The initial design \mathbf{X} and the initial family of sets of constant attributes C are input parameters. All information concerning the number of choice situations, the number of alternatives per choice situation and the profile strength is contained within \mathbf{X} and C . The design \mathbf{X} is composed of the design matrices \mathbf{X}_s corresponding to each choice situation s , with $1 \leq s \leq S$. Similarly, C is composed of the sets of attributes C_s that are held constant in each choice situation s , with $1 \leq s \leq S$. For example, the set $C_1 = \{1, 2\}$ indicates that attributes 1 and 2 are held constant in choice situation 1. The function \mathcal{D} calculates the value of the optimality criterion of a design. This value corresponds to the \mathcal{D}_0 -optimality criterion when generating UN designs, to the \mathcal{D}_p -optimality criterion when generating locally optimal designs, and to the \mathcal{D}_B -optimality criterion when generating Bayesian designs. The variable \mathcal{D}^0 stores the value of the optimality criterion before the optimization process begins.

Algorithm 1: Integrated algorithm

Input: The initial design \mathbf{X} and the initial family of sets of constant attributes C

```

1  $\mathcal{D}^0 \leftarrow \mathcal{D}(\mathbf{X})$ 
2 for  $s \leftarrow 1$  to  $S$  do
3    $\text{improve\_constant\_attributes}(\mathbf{X}, C, s)$  // See Algorithm 3
4    $\text{improve\_varying\_attributes}(\mathbf{X}, C, s)$  // See Algorithm 2
5 if  $\mathcal{D}(\mathbf{X}) > \mathcal{D}^0$  then
6    $\text{integrated\_algorithm}(\mathbf{X})$ 

```

The varying attributes within a choice situation are improved using a modified version of the coordinate-exchange algorithm proposed by [Meyer and Nachtsheim \(1995\)](#). This algorithm, which was introduced in the SC design literature by [Kessels et al. \(2009\)](#), iteratively exchanges the level of each attribute for the one that produces the best value of the optimality criterion. It applies a first-improvement strategy, which means that exchanges that lead to an improvement are performed as soon as they are found. The design \mathbf{X} , the family of sets of constant attributes C and the number of the choice situation s are input parameters of the coordinate-exchange algorithm. The pseudocode of this procedure is shown in [Algorithm 2](#). In this algorithm, the number of levels of attribute f is denoted by L_f and the levels themselves are written as $1, 2, \dots, L_f$. The variable $\mathbf{X}_{s,j,f}$ represents the level of attribute f in alternative j of choice situation s . The variable \mathcal{D}^* stores the best value of the optimality criterion found during the execution of [Algorithm 2](#). An illustrative explanation of the coordinate-exchange algorithm for quantitative attributes can be found in [Goos and Jones \(2011\)](#).

The set C_s of attributes that are held constant in choice situation s is improved one attribute at a time. At each iteration, one of the attributes is removed from the set (i.e., it is no longer considered constant) and treated as a varying attribute. Next, the algorithm evaluates the quality of the designs obtained by fixing one of the remaining attributes. All varying attributes are considered, and the one that produces the best value of the optimality criterion is added to the set of constant

Algorithm 2: Improving the levels of the varying factors (improve_varying_attributes)

Input: The design X , the family of sets of constant attributes C and the number of the choice situation, s , under consideration

```

1  $\mathcal{D}^0 \leftarrow \mathcal{D}(X)$ 
2 for  $j \leftarrow 1$  to  $J$  do
3   for  $f \leftarrow 1$  to  $F$  do
4     if  $f \notin C_s$  then
5        $\mathcal{D}^* \leftarrow \mathcal{D}(X)$ 
6        $l^* \leftarrow X_{s,j,f}$ 
7       for  $l \leftarrow 1$  to  $L_f$  do
8          $X_{s,j,f} \leftarrow l$ 
9         if  $\mathcal{D}(X) > \mathcal{D}^*$  then
10            $\mathcal{D}^* \leftarrow \mathcal{D}(X)$ 
11            $l^* \leftarrow l$ 
12        $X_{s,j,f} \leftarrow l^*$ 
13 if  $\mathcal{D}(X) > \mathcal{D}^0$  then
14   improve_varying_attributes( $X, C, s$ )

```

attributes. The algorithm finishes its execution when all the elements in the initial set of constant attributes have been evaluated for a potential exchange.

The pseudocode of the procedure to improve the set of constant attributes is shown in Algorithm 3. The goal of this algorithm is to find the best possible update of the family C of the sets of constant attributes C_1, \dots, C_S . The updated version of the family C is denoted by $C^* = \{C_1^*, \dots, C_S^*\}$. When an element c is removed from a set of constant attributes C_s , an additional improvement procedure is executed (line 6 in the pseudocode) to make the levels of the attribute different. This is done by applying the coordinate-exchange algorithm only to column c of the design matrix X_s (by executing lines 5 to 12 in Algorithm 2). Every time a varying attribute is considered for addition to C_s , it is necessary to store the levels for each alternative in a temporary vector T (line 9 in the pseudocode). This is because, once an attribute is considered fixed and has been added to C_s , it is set to a constant level (line 11 in the pseudocode). The attribute level assigned does not affect the

Algorithm 3: Improving the set of constant attributes (improve_constant_attributes)

Input: The design X , the family of sets of constant attributes C and the number of the choice situation, s , considered

```

1  $\mathcal{D}^* \leftarrow \mathcal{D}(X)$ 
2  $C^* \leftarrow C$ 
3 for  $c \in C_s$  do
4    $c^* \leftarrow c$ 
5    $C_s^* \leftarrow C_s^* \setminus \{c\}$ 
6   improve_varying_attributes( $X, C^*, s, c$ ) // Improve only the  $c$ -th column
7   for  $f \leftarrow 1$  to  $F$  do
8     if  $f \notin C_s^*$  then
9        $T \leftarrow X_{s,1:j,f}$ 
10       $C_s^* \leftarrow C_s^* \cup \{f\}$ 
11       $X_{s,1:j,f} \leftarrow \mathbf{1}_j$ 
12      if  $\mathcal{D}(X) > \mathcal{D}^*$  then
13         $\mathcal{D}^* \leftarrow \mathcal{D}(X)$ 
14         $c^* \leftarrow f$ 
15         $C_s^* \leftarrow C_s^* \setminus \{f\}$ 
16         $X_{s,1:j,f} \leftarrow T$ 
17    $C_s^* \leftarrow C_s^* \cup \{c^*\}$ 
18    $X_{s,1:j,c^*} \leftarrow \mathbf{1}_j$ 
19  $C \leftarrow C^*$ 

```

estimation of the MNL model or the quality of the design. However, for consistency purposes, the algorithm always sets the attribute to its first level. The original levels of the attribute are restored (line 16 in the pseudocode) after the quality of the resulting design has been evaluated.

As shown in [Algorithm 1](#), [Algorithms 3](#) and [2](#) are executed sequentially in order to improve each choice situation of the design. Improving the set of constant attributes first, followed by improving the varying attributes, leads to the best algorithm performance. During the development phase of the algorithm, we observed that the choice of constant attributes has the largest impact on the quality of the design. Therefore, the algorithm generates better designs by making this decision first. We also tried other alternative improvement strategies. None of them, despite being more complex and computationally more demanding, outperformed the simple algorithm described. For example, we tested a more exhaustive improvement procedure for the set of constant attributes. Instead of improving one constant attribute at a time, we tried improving various constant attributes at a time. We also tried improving the levels of all varying attributes every time a new set of constant attributes is evaluated. In that way, the final set of constant attributes would be chosen considering the best configuration of the varying attributes. However, none of these improvement strategies generated designs with a better quality.

We recommend the integrated algorithm to be executed using a multi-start strategy. In other words, we suggest executing the algorithm several times, starting from different initial random designs, and selecting the best design generated (the one that optimizes the value of the optimality criterion). The multi-start strategy is common to the algorithms described by [Huber and Zwerina \(1996\)](#), [Sándor and Wedel \(2001\)](#), [Kuhfeld and Tobias \(2005\)](#), [Rose et al. \(2008\)](#), [Kessels et al. \(2009\)](#), [Bliemer and Rose \(2010\)](#) and [Quan et al. \(2011\)](#). This strategy allows the algorithm to generate designs with a very good quality in a very short execution time.

When testing the algorithm on the UN benchmark instances described by [Großmann et al. \(2009\)](#), we noticed that, while it outperforms the algorithms of [Kessels et al. \(2011b, 2015\)](#), the algorithm is not able to find the optimal designs for some of the instances; not even when using a large number of restarts. The integrated algorithm generates designs that are as efficient as those suggested by [Großmann et al. \(2009\)](#) for 28 out of 50 instances. For the other 22 instances, the integrated algorithm produces designs that are slightly less efficient (at most 0.44% less efficient) than the optimal designs. The results of this computational experiment are described in depth in [Section 6.1](#) and shown in [Table 1](#). The limitation of the integrated algorithm is due to the fact that each component of the algorithm improves the design by making locally optimal decisions. Since the algorithm improves one choice situation at a time, it is not able to evaluate the repercussions that these simple improvements have on the structure (and, hence, the quality) of the entire design. This is especially so when selecting the set of constant attributes for each choice situation, which is the decision with the largest impact.

4. Extensive integrated algorithm

In order to overcome the limitation shown by the integrated algorithm, we propose an extra procedure that is executed on top of it. This additional procedure evaluates the effects that the selection of a set of constant attributes (for a particular choice situation) has on the entire design. This procedure iterates over the choice situations of the design and executes an algorithm similar to the one that improves the set of constant attributes of a choice situation (explained in [Section 3](#) and shown in [Algorithm 3](#)). Every time a new configuration of constant attributes is evaluated in a certain choice situation, the algorithm executes the integrated algorithm for all other choice situations too in order to propagate the effects of this new configuration. This leads to an exponential increase in complexity compared to the integrated algorithm. However, this intensive approach performs a wider and more thorough exploration of the design space and generates better designs. Due to this characteristic, we called this algorithm the *extensive integrated algorithm*.

The pseudocode of this new extensive algorithm is shown in [Algorithm 4](#). Like [Algorithm 3](#), [Algorithm 4](#) aims at finding the best set of constant attributes. At each iteration, the algorithm evaluates the replacement of one element in the set of constant attributes by a varying attribute. Nonetheless, there are two important differences between [Algorithms 3](#) and [4](#). First, in order to evaluate the quality of a replacement, [Algorithm 4](#) creates extra copies X^\dagger and C^\dagger of X and C , respectively (lines 8 and 9 of the pseudocode). This is because, once a replacement has been carried out, the integrated algorithm is executed to improve the slightly modified design (line 14 of the pseudocode). By doing so, the effects of the change in the set of constant attributes of a given choice situation are propagated to the other choice situations of the design. The resulting design is the best one generated by the integrated algorithm if the new set of constant attributes evaluated is considered part of the initial structure. If this new design has a better quality than the best design found so far, it is stored in variables X^* and C^* (lines 15, 16 and 17 of the pseudocode). Note that it is possible that the resulting design has a very different structure compared to the slightly modified initial design. The second difference between [Algorithms 3](#) and [4](#) is related to the number of exchanges evaluated by the algorithms when improving a choice situation. Observe that, in [Algorithm 4](#), the original design is updated if a better design is found after evaluating the exchange of one element in the set of constant attributes (lines 18, 19 and 20 of the pseudocode). In this event, because the entire family of constant attributes potentially experienced a major change due to the nested execution of the integrated algorithm, it is necessary to restart the evaluation of the set C_s (the loop starting in line 3 of the pseudocode).

Algorithm 4: Extensive algorithm**Input:** The initial design \mathbf{X} and the initial family of sets of constant attributes C

```

1  $\mathcal{D}^0 \leftarrow \mathcal{D}(\mathbf{X})$ 
2 for  $s \leftarrow 1$  to  $S$  do
3   for  $c \in C_s$  do
4      $\mathbf{X}^* \leftarrow \mathbf{X}$ 
5      $C^* \leftarrow C$ 
6     for  $f \leftarrow 1$  to  $F$  do
7       if  $f \notin C_s$  then
8          $\mathbf{X}^\dagger \leftarrow \mathbf{X}$ 
9          $C^\dagger \leftarrow C$ 
10         $C_s^\dagger \leftarrow C_s^\dagger \setminus \{c\}$ 
11        improve_varying_attributes( $\mathbf{X}^\dagger, C^\dagger, s, c$ ) // Improve only the  $c$ -th column
12         $C_s^\dagger \leftarrow C_s^\dagger \cup \{f\}$ 
13         $\mathbf{X}_{s,1:j,f}^\dagger \leftarrow \mathbf{1}_j$ 
14        integrated_algorithm( $\mathbf{X}^\dagger, C^\dagger$ )
15        if  $\mathcal{D}(\mathbf{X}^\dagger) > \mathcal{D}(\mathbf{X}^*)$  then
16           $\mathbf{X}^* \leftarrow \mathbf{X}^\dagger$ 
17           $C^* \leftarrow C^\dagger$ 
18        if  $\mathbf{X} \neq \mathbf{X}^* \vee C \neq C^*$  then
19           $\mathbf{X} \leftarrow \mathbf{X}^*$ 
20           $C \leftarrow C^*$ 
21 if  $\mathcal{D}(\mathbf{X}) > \mathcal{D}^0$  then
22   extensive_algorithm( $\mathbf{X}, C$ )

```

5. Implementation details

Both algorithms, the integrated algorithm from Section 3 and the extensive algorithm from Section 4, were implemented using C++. In order to speed up the execution of the algorithms, we paid special attention to the calculation of the optimality criteria. The matrix \mathbf{X}_s of each choice situation s is separately stored in memory. By doing so, it is possible to implement update procedures that avoid recalculating the entire information matrix every time a choice situation is modified. When a new design \mathbf{X}^* is generated by modifying choice situation s of a design \mathbf{X} , it is possible to calculate the corresponding information matrix by using the formula

$$\mathbf{M}(\mathbf{X}^*, \boldsymbol{\beta}) = \mathbf{M}(\mathbf{X}, \boldsymbol{\beta}) - \mathbf{f}'(\mathbf{X}_s)(\mathbf{P}_s - \mathbf{p}_s \mathbf{p}_s') \mathbf{f}(\mathbf{X}_s) + \mathbf{f}'(\mathbf{X}_s^*)(\mathbf{P}_s^* - \mathbf{p}_s^* \mathbf{p}_s^{*'}) \mathbf{f}(\mathbf{X}_s^*), \quad (8)$$

where \mathbf{P}_s^* and \mathbf{p}_s^* are calculated using the matrix \mathbf{X}_s^* of choice situation s in the new design \mathbf{X}^* . The new information matrix is calculated from the initial one by subtracting the part of it corresponding to the initial choice situation and adding the term corresponding to the new one (see Eqs. (3) and (4)). This strategy increases the memory used by the algorithms, but considerably reduces their execution times.

In order to calculate the K -dimensional integral in Eq. (5) (related to the prior distribution of the parameter vector $\boldsymbol{\beta}$) when generating \mathcal{D}_B -optimal designs, the algorithms use the integration method proposed by Gotwalt et al. (2009) and Gotwalt (2010). This method is based on the radial-spherical integration rule proposed by Monahan and Genz (1997). It transforms the integral in Eq. (5) into a radial component and a spherical surface component. The radial integration is performed using the generalized Gauss-Laguerre quadrature while the spherical integration is performed using a randomly rotated Mysovskikh (1980) extended simplex quadrature. Bliemer et al. (2008) and Yu et al. (2010) showed that quadrature methods usually outperform other methods for evaluating the Bayesian optimality criterion. Contrary to Monte Carlo sampling, quadrature methods estimate the integral by evaluating the values for the \mathcal{D}_B -optimality criterion using a small systematic sample of parameter vectors $\boldsymbol{\beta}$ from the prior distribution. In order to maximize the speed of our algorithm, the calculations related to each vector $\boldsymbol{\beta}$ in the sample are stored separately. This allows us to implement the update formula in Eq. (8) for each sampled vector $\boldsymbol{\beta}$.

6. Computational experiments

In this section, we describe the results of several computational experiments we carried out with two main objectives. The first objective was to assess the performance of the algorithms under different experimental conditions. The second objective was to analyse the effects that these conditions have on the structure of the designs generated. Since there are many factors that could influence the algorithms' behaviour and the final structure of the designs, we executed four different experiments, each with a specific purpose. The first experiment was carried out in order to (i) obtain a general overview of the algorithms' performance and (ii) perform a first comparison to the latest version of two-stage algorithm (Kessels et al., 2015) and to the analytical generation method of Großmann et al. (2009). The purpose of the second experiment was to study the influence of the number of choice situations, the number of alternatives per choice situation, the profile strength and the attributes' heterogeneity in terms of the number of levels on the performance of the algorithms. The third experiment was carried out to study the effect of the parameter vector β when generating locally optimal designs. Finally, the purpose of the fourth experiment was to study the influence of the variance-covariance matrix of the prior distribution when generating Bayesian optimal designs. The experiments were executed using a computer with a 2.93 GHz Intel Core i7 processor. The two-stage algorithm of Kessels et al. (2015) and the integrated algorithm were executed using 1000 restarts, and the extensive (integrated) algorithm was executed using 25 restarts. All designs generated in this section assume categorical attributes that are effects-type coded.

6.1. General overview and preliminary comparison

In order to obtain a general overview of the algorithms' performance, we ran the two-stage algorithm, the integrated algorithm and the extensive algorithm to generate UN designs for the 50 instances listed by Großmann et al. (2009). These instances involve different SC experiments with two alternatives per choice situation. The results are shown in Table 1. The first five columns show information about the instances: the name, the number of attributes (F), the number of levels for each attribute, the profile strength (G) and the number of choice situations (S). The other columns show the \mathcal{D}_0 -efficiencies of the designs generated by each algorithm relative to the optimal designs analytically constructed by Großmann et al. (2009). For the two-stage algorithm, only the maximum \mathcal{D}_0 -efficiencies (corresponding to the best designs found during the 1000 restarts) are shown. For the integrated and the extensive algorithm, the minimum and the average \mathcal{D}_0 -efficiencies are also displayed, along with the average execution time of each restart. We do not report the average execution time of the two-stage algorithm because it has been implemented as a script in the statistical software JMP. This script takes a considerably longer time to execute than the integrated algorithm and its extensive version. However, this is not due to a greater complexity but due to performance differences between JMP's scripting language and C++.

The best designs generated by the integrated algorithm and the extensive algorithm (whose efficiencies are shown in Table 1's columns labelled "Max.") consistently outperform the designs generated by the two-stage algorithm. Moreover, the best designs produced by the integrated algorithm have \mathcal{D}_0 -efficiencies that are very close to 100%. The best design with the lowest quality produced by the integrated algorithm, corresponding to instance PP40, is only 0.44% less efficient than the optimal design constructed by Großmann et al. (2009). Despite the fact that its execution time is very short, the integrated algorithm found optimal designs for 28 of the 50 instances. Its longest average execution time is less than one second. On the contrary, the extensive algorithm generates optimal designs for all instances, except for instance PP40. To our knowledge, the extensive algorithm is the first algorithm that is able to match the quality of the designs analytically derived. This effectiveness comes, however, at a price: the execution time of the extensive algorithm is several orders of magnitude larger than the execution time of the simple integrated algorithm.

The two-stage algorithm generates designs with large \mathcal{D}_0 -efficiencies for many of the experiments in the benchmark set. However, the \mathcal{D}_0 -efficiencies of the best designs generated by the integrated algorithm are usually larger. The difference in \mathcal{D}_0 -efficiencies between the two algorithms is larger when the number of choice situations is small and all but one or two attributes have the same number of levels. For example, for instance PP14, involving three two-level attributes and one five-level attribute, the design generated by the two-stage algorithm is 8.88% less \mathcal{D}_0 -efficient than the one generated by the integrated algorithm. This difference in \mathcal{D}_0 -efficiency could be interpreted in two ways. First, the design generated by the integrated algorithm is able to collect close to 9% more information than the design constructed by the two-stage algorithm. Second, the design generated by the integrated algorithm can gather as much information as the design constructed by the two-stage algorithm with 9% fewer respondents. For most of the instances, even the minimum \mathcal{D}_0 -efficiencies achieved by the integrated and the extensive algorithm are larger than the maximum \mathcal{D}_0 -efficiencies achieved by the two-stage algorithm. This shows the benefits of simultaneously optimizing the sets of constant attributes and the levels of the varying attributes when generating an optimal design for a SC experiment with partial profiles.

For some instances, the integrated algorithm and the extensive algorithm generate designs with very different structures than those proposed by Großmann et al. (2009). One such instance is PP23, involving three two-level attributes and two four-level attributes. Table 2 compares the designs generated for this instance by the different approaches. The first column of the table shows the number of the choice situation. Each subsequent column groups the choice situations that constitute one of the designs compared. Each element of the column therefore shows an alternative (in a choice situation) and its corresponding attribute levels. The attributes that are held constant in each choice situation

Table 1

Performance comparison of the two-stage, the integrated and the extensive algorithm using a benchmark set of 50 instances.

Instance information					Two-stage	Integrated				Extensive			
					\mathcal{D}_0 -eff.	\mathcal{D}_0 -efficiency			Time (s)	\mathcal{D}_0 -efficiency			Time (s)
Name	F	Levels	G	S	Max.	Min.	Avg.	Max.	Avg.	Min.	Avg.	Max.	Avg.
PP01	4	2,3,3,3	3	42	99.59	99.70	99.86	99.97	0.02	99.96	99.97	100.00	4.03
PP02	4	2,2,3,3	2	18	90.31	97.84	99.29	100.00	0.01	100.00	100.00	100.00	0.75
PP03	4	2,2,3,3	3	12	99.07	95.78	98.30	100.00	0.01	99.53	99.81	100.00	0.24
PP04	4	2,2,4,4	2	16	88.45	96.31	98.39	100.00	0.01	99.20	99.70	100.00	1.02
PP05	4	2,2,4,4	3	24	98.99	98.82	99.48	99.91	0.02	99.72	99.85	100.00	1.54
PP06	4	2,2,5,5	2	50	97.24	99.39	99.73	99.92	0.06	99.90	99.96	100.00	15.61
PP07	4	2,2,5,5	3	40	99.90	99.39	99.68	99.90	0.05	99.85	99.92	100.00	7.25
PP08	4	3,3,4,4	2	60	97.81	99.61	99.82	99.95	0.06	99.95	99.99	100.00	20.97
PP09	4	3,3,5,5	2	90	97.04	99.78	99.88	99.97	0.13	99.97	99.99	100.00	65.71
PP10	4	2,2,2,3	2	30	99.13	99.46	99.86	100.00	0.00	99.72	99.99	100.00	1.47
PP11	4	2,2,2,3	3	36	100.00	99.11	99.90	100.00	0.01	99.90	99.99	100.00	1.68
PP12	4	2,2,2,4	2	12	93.86	95.65	98.17	100.00	0.00	100.00	100.00	100.00	0.40
PP13	4	2,2,2,4	3	72	100.00	99.81	99.96	100.00	0.02	99.98	99.90	100.00	8.01
PP14	4	2,2,2,5	2	60	91.12	99.79	99.93	100.00	0.03	100.00	100.00	100.00	10.82
PP15	4	3,3,3,4	2	54	99.72	99.59	99.81	100.00	0.04	99.92	99.98	100.00	12.69
PP16	5	2,3,3,3,3	3	36	99.58	99.35	99.66	100.00	0.04	99.86	99.91	100.00	14.46
PP17	5	2,2,3,3,3	2	24	97.70	98.23	99.34	100.00	0.01	99.74	99.95	100.00	4.17
PP18	5	2,2,3,3,3	3	96	99.40	99.93	99.97	99.99	0.09	99.99	99.99	100.00	69.93
PP19	5	2,2,4,4,4	2	44	98.78	99.19	99.58	99.89	0.07	99.86	99.93	100.00	25.72
PP20	5	2,2,5,5,5	2	70	98.14	99.52	99.73	99.91	0.22	99.93	99.96	100.00	102.38
PP21	5	2,2,2,3,3	2	42	97.38	99.47	99.82	100.00	0.02	99.93	99.99	100.00	9.57
PP22	5	2,2,2,3,3	3	28	97.65	99.07	99.63	99.93	0.02	99.83	99.92	100.00	5.53
PP23	5	2,2,2,4,4	2	18	93.66	96.02	98.59	100.00	0.02	99.29	99.84	100.00	2.75
PP24	5	2,2,2,4,4	3	24	98.18	98.08	99.06	100.00	0.04	99.48	99.83	100.00	5.97
PP25	5	3,3,3,4,4	2	72	99.55	99.64	99.82	100.00	0.14	99.96	99.98	100.00	67.33
PP26	5	3,3,3,4,4	3	96	99.63	99.86	99.92	99.97	0.25	99.97	99.99	100.00	173.55
PP27	5	2,2,2,2,3	2	36	98.70	99.23	99.79	100.00	0.02	99.77	99.98	100.00	5.04
PP28	5	2,2,2,2,3	3	24	98.84	97.08	99.36	100.00	0.01	99.28	99.66	100.00	3.35
PP29	5	2,2,2,2,4	2	28	96.74	98.71	99.59	100.00	0.02	99.78	99.99	100.00	4.02
PP30	5	2,2,2,2,4	3	24	99.60	97.90	99.31	100.00	0.02	99.43	99.75	100.00	3.62
PP31	5	2,2,2,2,5	2	40	94.87	99.11	99.70	100.00	0.03	100.00	100.00	100.00	11.90
PP32	5	2,2,2,2,5	3	40	97.75	98.55	99.62	100.00	0.04	99.65	99.86	100.00	12.18
PP33	6	2,2,3,3,3,3	2	30	98.19	98.10	99.32	100.00	0.05	99.72	99.91	100.00	12.91
PP34	6	2,2,4,4,4,4	2	28	93.75	96.60	98.44	100.00	0.10	99.32	99.89	100.00	22.13
PP35	6	2,2,5,5,5,5	2	90	98.57	99.53	99.71	99.90	0.69	99.92	99.97	100.00	485.10
PP36	6	3,3,4,4,4,4	2	96	99.67	99.70	99.81	99.93	0.44	99.97	99.99	100.00	350.39
PP37	6	2,2,2,3,3,3	2	54	99.27	99.57	99.82	100.00	0.07	99.94	99.98	100.00	36.09
PP38	6	2,2,2,3,3,3	3	36	99.15	99.30	99.64	99.92	0.07	99.86	99.92	100.00	24.17
PP39	6	2,2,2,4,4,4	2	48	98.17	99.18	99.59	99.90	0.11	99.87	99.95	100.00	49.42
PP40	6	2,2,2,4,4,4	3	32	98.13	98.49	99.08	99.56	0.11	99.51	99.67	99.90	35.89
PP41	6	2,2,2,5,5,5	3	100	97.81	99.79	99.88	99.94	0.72	99.95	99.97	100.00	625.57
PP42	6	2,2,2,2,3,3	2	24	97.03	97.61	99.34	100.00	0.02	100.00	100.00	100.00	4.79
PP43	6	2,2,2,2,3,3	3	32	99.03	99.17	99.60	99.87	0.04	99.83	99.91	100.00	16.32
PP44	6	2,2,2,2,4,4	2	20	93.19	96.59	98.69	100.00	0.03	99.52	99.90	100.00	5.83
PP45	6	2,2,2,2,4,4	3	80	98.09	99.83	99.92	99.98	0.18	99.97	99.99	100.00	133.49
PP46	6	2,2,2,2,5,5	2	60	96.83	99.46	99.73	99.94	0.14	99.92	99.95	100.00	83.89
PP47	6	2,2,2,2,5,5	3	40	96.37	98.78	99.40	99.83	0.15	99.61	99.87	100.00	59.86
PP48	6	3,3,3,3,4,4	2	84	99.60	99.67	99.81	99.96	0.25	99.96	99.97	100.00	187.71
PP49	6	2,2,2,2,2,4	2	80	97.81	99.73	99.93	100.00	0.08	100.00	100.00	100.00	64.22
PP50	6	2,2,2,2,2,5	2	90	96.31	99.81	99.93	100.00	0.12	100.00	100.00	100.00	97.19
Average					97.51	98.80	99.52	99.96	0.10	99.80	99.93	99.99	59.37

are marked with an asterisk. The design analytically generated by Großmann et al. (2009) involves 12 choice situations where the three two-level attributes are held constant, and 6 choice situations where one two-level attribute and the two four-level attributes are held constant. The design generated by the integrated algorithm involves 6 choice situations where the three two-level attributes are held constant, and 12 choice situations where two two-level attributes and one four-level attribute are held constant. The design generated by the extensive algorithm involves 8 choice situations where the three two-level attributes are held constant, 8 choice situations where two two-level attributes and one four-level attribute are held constant, and 2 choice situations where one two-level attribute and the two four-level attributes are held constant. The existence of such designs, different from the ones generated by Großmann et al. (2009) but equally efficient, was hitherto unknown. The design generated by the two-stage algorithm is 6.34% less \mathcal{D}_0 -

Table 2

UN designs for instance PP23 involving 18 choice situations, 2 alternatives per choice situation, 3 two-level attributes, 2 four-level attributes and a profile strength of 2.

CS	Großmann et al.					Integrated					Extensive					Two-stage				
	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5
1	*	*	*	1	1	*	*	*	4	1	*	*	*	2	4	*	*	*	2	1
	*	*	*	2	2	*	*	*	1	3	*	*	*	3	1	*	*	*	3	3
2	*	*	*	1	1	*	*	*	4	4	*	*	*	1	4	*	*	*	2	3
	*	*	*	3	3	*	*	*	3	3	*	*	*	2	2	*	*	*	4	1
3	*	*	*	1	1	*	*	*	1	3	*	*	*	3	4	*	*	*	3	1
	*	*	*	4	4	*	*	*	2	2	*	*	*	2	1	*	*	*	4	2
4	*	*	*	2	2	*	*	*	4	3	*	*	*	4	4	*	*	*	3	2
	*	*	*	3	3	*	*	*	1	1	*	*	*	3	3	*	*	*	4	3
5	*	*	*	2	2	*	*	*	2	3	*	*	*	4	3	*	*	1	*	2
	*	*	*	4	4	*	*	*	1	2	*	*	*	1	1	*	*	2	*	4
6	*	*	*	3	3	*	*	*	3	4	*	*	*	4	1	*	*	1	*	4
	*	*	*	4	4	*	*	*	4	3	*	*	*	1	3	*	*	2	*	2
7	*	*	*	1	2	*	*	1	*	2	*	*	*	3	4	*	*	1	1	*
	*	*	*	2	1	*	*	2	*	1	*	*	*	4	3	*	*	2	4	*
8	*	*	*	1	3	*	*	2	*	2	*	*	*	2	4	*	*	1	4	*
	*	*	*	3	1	*	*	1	*	1	*	*	*	1	2	*	*	2	1	*
9	*	*	*	1	4	*	*	2	4	*	*	*	2	*	2	*	1	*	*	3
	*	*	*	4	1	*	*	1	2	*	*	*	1	*	3	*	2	*	*	4
10	*	*	*	2	3	*	*	1	4	*	*	*	2	*	3	*	1	*	*	4
	*	*	*	3	2	*	*	2	2	*	*	*	1	*	2	*	2	*	*	3
11	*	*	*	2	4	*	2	*	*	4	*	*	2	3	*	*	1	*	1	*
	*	*	*	4	2	*	1	*	*	2	*	*	1	1	*	*	2	*	3	*
12	*	*	*	3	4	*	2	*	*	2	*	*	1	3	*	*	1	*	3	*
	*	*	*	4	3	*	1	*	*	4	*	*	2	1	*	*	2	*	1	*
13	*	1	1	*	*	*	2	*	1	*	*	1	*	*	2	1	*	*	*	4
	*	2	2	*	*	*	1	*	3	*	*	2	*	*	1	2	*	*	*	1
14	*	1	2	*	*	*	2	*	3	*	*	1	*	*	1	1	*	*	1	*
	*	2	1	*	*	*	1	*	1	*	*	2	*	*	2	2	*	*	2	*
15	1	*	1	*	*	2	*	*	*	1	2	*	*	2	*	1	*	*	2	*
	2	*	2	*	*	1	*	*	*	4	1	*	*	4	*	2	*	*	1	*
16	1	*	2	*	*	2	*	*	*	4	1	*	*	2	*	1	*	2	*	*
	2	*	1	*	*	1	*	*	*	1	2	*	*	4	*	2	*	1	*	*
17	1	1	*	*	*	2	*	*	3	*	2	2	*	*	*	1	1	*	*	*
	2	2	*	*	*	1	*	*	2	*	1	1	*	*	*	2	2	*	*	*
18	1	2	*	*	*	1	*	*	3	*	2	1	*	*	*	1	2	*	*	*
	2	1	*	*	*	2	*	*	2	*	1	2	*	*	*	2	1	*	*	*
\mathcal{D}_0 -eff.	100.00					100.00					100.00					93.66				

efficient than the other designs, and appears to have too few choice situations where the three two-level attributes are held constant.

For some other instances, the designs produced by the integrated algorithm and the extensive algorithm are similar in structure to those produced by Großmann et al. (2009). Table 3 illustrates this for instance PP24, also involving three two-level attributes and two four-level attributes, where the optimal designs generated by these approaches only hold the two-level attributes constant. The design generated by the two-stage algorithm has a different structure (it also involves choice situations in which one four-level attribute is held constant) and is 1.82% less \mathcal{D}_0 -efficient than the optimal design.

6.2. Influence of the experimental setup

The benchmark set used in the previous experiment includes a considerable number of instances. However, there are several experimental setups that are not considered. For example, the set does not include SC experiments with more than two alternatives per choice situation. Since the algorithms that we propose are very general, a more structured set of instances that allows the evaluation of the algorithm in a wider range of situations was needed. Therefore, we identified the parameters that define a SC experiment and ran a full factorial experiment with the most interesting levels for these parameters. The parameters considered are the number of choice situations (S), the number of alternatives per choice situation (J), the profile strength (G) and the attributes' heterogeneity in terms of the number of levels. The parameters (of the

Table 3

UN designs for instance PP24 involving 24 choice situations, 2 alternatives per choice situation, 3 two-level attributes, 2 four-level attributes and a profile strength of 3.

CS	Großmann et al.					Integrated					Extensive					Two-stage				
	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5	1	2	3	4	5
1	*	*	1	2	2	*	*	1	3	2	*	*	2	4	4	*	*	1	2	1
	*	*	2	4	4	*	*	2	1	4	*	*	1	2	2	*	*	2	4	4
2	*	*	1	2	4	*	*	2	3	1	*	*	1	4	4	*	*	1	2	4
	*	*	2	4	2	*	*	1	4	4	*	*	2	1	1	*	*	2	1	1
3	*	*	1	3	3	*	*	1	2	4	*	*	2	2	4	*	*	1	3	1
	*	*	2	4	4	*	*	2	3	2	*	*	1	1	1	*	*	2	1	3
4	*	*	1	3	4	*	*	2	4	1	*	*	2	4	2	*	*	1	4	2
	*	*	2	4	3	*	*	1	3	2	*	*	1	3	4	*	*	2	3	4
5	*	*	1	4	2	*	*	2	4	2	*	*	2	3	3	*	*	1	4	3
	*	*	2	2	4	*	*	1	2	3	*	*	1	1	2	*	*	2	1	4
6	*	*	1	4	3	*	*	2	3	4	*	*	1	3	1	*	*	1	4	4
	*	*	2	3	4	*	*	1	4	1	*	*	2	1	3	*	*	2	2	2
7	*	*	1	4	4	*	*	1	1	2	*	*	2	3	2	*	1	*	1	2
	*	*	2	2	2	*	*	2	2	3	*	*	1	2	3	*	2	*	3	1
8	*	*	1	4	4	*	*	2	2	2	*	*	2	2	1	*	1	*	2	4
	*	*	2	3	3	*	*	1	3	1	*	*	1	4	3	*	2	*	4	2
9	*	1	*	1	1	*	1	*	2	2	*	2	*	3	3	*	1	*	3	2
	*	2	*	4	4	*	2	*	1	4	*	1	*	1	4	*	2	*	1	1
10	*	1	*	1	4	*	2	*	3	2	*	2	*	3	4	*	1	*	3	3
	*	2	*	4	1	*	1	*	1	1	*	1	*	4	3	*	2	*	2	4
11	*	1	*	2	2	*	2	*	1	1	*	2	*	4	2	*	1	*	4	1
	*	2	*	3	3	*	1	*	3	4	*	1	*	2	1	*	2	*	2	3
12	*	1	*	2	3	*	1	*	4	4	*	2	*	1	3	*	1	*	4	3
	*	2	*	3	2	*	2	*	1	3	*	1	*	3	2	*	2	*	3	2
13	*	1	*	3	3	*	2	*	2	2	*	2	*	2	2	*	1	1	1	*
	*	2	*	2	2	*	1	*	3	1	*	1	*	1	4	*	2	2	4	*
14	*	1	*	4	4	*	1	*	2	3	*	1	*	2	3	*	1	2	2	*
	*	2	*	1	1	*	2	*	3	1	*	2	*	4	1	*	2	1	1	*
15	*	1	*	3	2	*	2	*	4	3	*	2	*	1	1	1	*	*	1	2
	*	2	*	2	3	*	1	*	1	2	*	1	*	4	2	2	*	*	2	1
16	*	1	*	4	1	*	2	*	2	4	*	2	*	2	4	1	*	*	2	2
	*	2	*	1	4	*	1	*	1	3	*	1	*	3	1	2	*	*	1	3
17	1	*	*	1	1	1	*	*	1	2	2	*	*	2	1	1	*	*	2	3
	2	*	*	2	2	2	*	*	3	4	1	*	*	1	2	2	*	*	3	2
18	1	*	*	1	1	1	*	*	2	1	2	*	*	2	2	1	*	*	3	3
	2	*	*	3	3	2	*	*	4	3	1	*	*	3	3	2	*	*	1	1
19	1	*	*	1	2	1	*	*	4	4	1	*	*	3	4	1	*	*	3	4
	2	*	*	2	1	2	*	*	2	3	2	*	*	4	1	2	*	*	2	2
20	1	*	*	1	3	2	*	*	4	2	1	*	*	4	1	1	*	*	4	4
	2	*	*	3	1	1	*	*	3	3	2	*	*	1	3	2	*	*	3	3
21	1	*	*	2	1	2	*	*	1	3	2	*	*	4	4	1	*	2	*	1
	2	*	*	1	2	1	*	*	2	1	1	*	*	1	2	2	*	1	*	3
22	1	*	*	2	2	1	*	*	4	3	2	*	*	3	3	1	*	1	1	*
	2	*	*	1	1	2	*	*	2	1	1	*	*	2	4	2	*	2	4	*
23	1	*	*	3	1	1	*	*	1	4	2	*	*	3	2	1	1	*	*	1
	2	*	*	1	3	2	*	*	4	1	1	*	*	4	1	2	2	*	*	4
24	1	*	*	3	3	2	*	*	1	4	2	*	*	1	4	1	2	*	*	3
	2	*	*	1	1	1	*	*	4	3	1	*	*	2	3	2	1	*	*	4
D₀-eff.	100.00					100.00					100.00					98.18				

computational experiment) and their levels are shown in Table 4. Observe that the attributes' heterogeneity is defined as a specific configuration of the numbers of attribute levels. The number of attributes was fixed to 6. It was not included as a factor in the study to limit the size of the experiment. The total number of instances considered in the factorial experiment was $2^3 \times 3 = 24$.

Table 4

Parameters and levels tested in Section 6.2 to study the influence of the SC experiment's characteristics on the performance of the two-stage, the integrated and the extensive algorithm.

Factor	Levels
S	15, 30
J	2, 3
G	3, 4
Attributes' heterogeneity	(3, 3, 3, 3, 3, 3), (2, 2, 2, 4, 4, 4), (2, 2, 3, 3, 4, 4)

Table 5

Performance comparison of the two-stage, the integrated and the extensive algorithm for the 24 instances in the factorial experiment described in Section 6.2.

Factors				Two-stage	Integrated		Extensive	
S	J	G	Num. of levels	\mathcal{D}_0 -eff.	\mathcal{D}_0 -eff.	Time (s)	\mathcal{D}_0 -eff. ^a	Time (s)
15	2	4	3,3,3,3,3	99.18	99.34	0.04	100.00	6.77
15	2	4	2,2,2,4,4	95.43	99.84	0.05	100.00	8.15
15	2	4	2,2,3,3,4	98.02	99.19	0.05	100.00	7.65
15	2	3	3,3,3,3,3	95.72	99.72	0.04	100.00	8.91
15	2	3	2,2,2,4,4	95.03	98.92	0.05	100.00	9.00
15	2	3	2,2,3,3,4	95.25	98.11	0.04	100.00	9.49
15	3	4	3,3,3,3,3	99.69	99.66	0.07	100.00	10.15
15	3	4	2,2,2,4,4	98.03	99.78	0.08	100.00	14.91
15	3	4	2,2,3,3,4	98.64	99.71	0.08	100.00	14.49
15	3	3	3,3,3,3,3	98.14	100.00	0.07	100.00	12.61
15	3	3	2,2,2,4,4	97.61	99.79	0.08	100.00	17.11
15	3	3	2,2,3,3,4	98.45	99.71	0.08	100.00	15.58
30	2	4	3,3,3,3,3	99.75	99.81	0.09	100.00	27.10
30	2	4	2,2,2,4,4	93.02	99.84	0.11	100.00	30.64
30	2	4	2,2,3,3,4	98.61	99.71	0.10	100.00	29.71
30	2	3	3,3,3,3,3	99.57	99.74	0.09	100.00	30.76
30	2	3	2,2,2,4,4	97.01	99.95	0.10	100.00	29.86
30	2	3	2,2,3,3,4	97.91	99.76	0.10	100.00	31.76
30	3	4	3,3,3,3,3	99.85	99.91	0.15	100.00	38.60
30	3	4	2,2,2,4,4	98.56	99.93	0.18	100.00	66.59
30	3	4	2,2,3,3,4	98.71	99.91	0.17	100.00	54.50
30	3	3	3,3,3,3,3	98.62	99.85	0.16	100.00	56.05
30	3	3	2,2,2,4,4	98.35	99.92	0.17	100.00	85.86
30	3	3	2,2,3,3,4	98.46	99.88	0.17	100.00	68.56
Average				97.82	99.67	0.10	100.00	28.53

^a Since the designs generated by the extensive algorithm were considered to be optimal, these efficiencies are set to 100%.

The two-stage algorithm, the integrated algorithm and the extensive algorithm were used to generate UN designs for each of the 24 instances in the factorial experiment. Since there exist no analytical results concerning optimal designs for these instances, the designs obtained by the extensive algorithm were considered optimal. The \mathcal{D}_0 -efficiencies of the designs generated by the two-stage and the integrated algorithm were therefore calculated relative to those produced by the extensive algorithm. The results are shown in Table 5.

The results confirm that the integrated algorithm and the extensive algorithm generate designs with a better quality than those produced by the two-stage algorithm. This is corroborated by three statistical tests:

- An analysis of variance (ANOVA) was conducted using the \mathcal{D}_0 -efficiency as a response and including the type of algorithm as a factor, in addition to the factors in Table 4. The type of algorithm was found to have a significant impact on the \mathcal{D}_0 -efficiency with a p -value smaller than 0.001.
- Both the sign test and the Wilcoxon signed-rank test found the average \mathcal{D}_0 -efficiency achieved by the integrated algorithm to be significantly larger than the average achieved by the two-stage algorithm (with p -values smaller than 0.001).

Additionally, the results in Table 5 show that the designs generated by the two-stage algorithm are particularly poor when $J = 2$, half of the attributes have 2 levels and the other half have 4 levels.

6.3. Influence of the prior parameter vector β

The parameter vector β is another factor that might influence the design generation process. This vector affects the calculation of the optimality criteria through the probabilities in Eq. (2). In the previous computational experiments, we set

Table 6

Factors and levels tested in Section 6.3 to study the influence of the prior parameter vector β on the performance of the two-stage, the integrated and the extensive algorithm.

Factor	Levels
Number of attributes with non-zero marginal utilities	2, 4, 6
Marginal utility of the least preferred level	-0.5, -1, -1.5, -2

Table 7

Performance comparison of the two-stage, the integrated and the extensive algorithm for 13 prior parameter vectors β .

β	Two-stage		Integrated		Extensive	
	\mathcal{D}_P -eff.	Num. const.	\mathcal{D}_P -eff.	Num. const.	\mathcal{D}_P -eff. ^a	Num. const.
0 0 0 0 0 0 0 0 0 0 0 0	99.42	5 5 5 5 5 5	99.42	6 6 4 4 5 5	100.00	5 5 6 6 4 4
-0.5 0 -0.5 0 0 0 0 0 0 0 0 0	98.35	5 5 5 5 5 5	99.50	5 7 4 5 5 4	100.00	6 6 4 5 4 5
-0.5 0 -0.5 0 -0.5 0 0 0 0 0 0 0	99.83	5 5 5 5 5 5	99.83	5 5 6 6 4 4	100.00	5 5 6 6 4 4
-0.5 0 -0.5 0 -0.5 0 -0.5 0 -0.5 0 -0.5 0	99.34	5 5 5 5 5 5	99.34	5 5 5 5 5 5	100.00	4 5 5 5 6 5
-1.0 0 -1.0 0 0 0 0 0 0 0 0 0	96.40	5 5 5 5 5 5	99.09	7 7 4 3 5 4	100.00	8 7 4 4 4 3
-1.0 0 -1.0 0 -1.0 0 -1.0 0 0 0 0 0	98.18	5 5 5 5 5 5	99.42	7 4 7 5 3 4	100.00	6 5 6 6 4 3
-1.0 0 -1.0 0 -1.0 0 -1.0 0 -1.0 0 -1.0 0	97.78	5 5 5 5 5 5	98.35	5 4 5 6 4 6	100.00	5 6 5 4 4 6
-1.5 0 -1.5 0 0 0 0 0 0 0 0 0	91.01	5 5 5 5 5 5	99.92	8 7 4 3 4 4	100.00	7 6 5 4 4 4
-1.5 0 -1.5 0 -1.5 0 -1.5 0 0 0 0 0	97.45	5 5 5 5 5 5	97.94	5 7 7 5 3 3	100.00	5 7 6 5 4 3
-1.5 0 -1.5 0 -1.5 0 -1.5 0 -1.5 0 -1.5 0	97.45	5 5 5 5 5 5	96.64	4 5 4 5 6 6	100.00	6 4 3 5 6 6
-2.0 0 -2.0 0 0 0 0 0 0 0 0 0	81.94	5 5 5 5 5 5	99.67	9 7 3 4 3 4	100.00	7 9 3 4 2 5
-2.0 0 -2.0 0 -2.0 0 -2.0 0 0 0 0 0	97.21	5 5 5 5 5 5	97.21	6 7 5 6 2 4	100.00	6 7 6 5 4 2
-2.0 0 -2.0 0 -2.0 0 -2.0 0 -2.0 0 -2.0 0	96.56	5 5 5 5 5 5	96.56	5 4 4 5 6 6	100.00	4 4 6 6 5 5
Average	96.22		98.68		100.00	

^a Since the designs generated by the extensive algorithm were considered to be optimal, these efficiencies are set to 100%.

$\beta = \mathbf{0}_K$ since various analytical results exist for that case. However, it is often more realistic to start from a different prior parameter vector. For example, it is known that people prefer short travel times over long ones, and low travel costs over high ones. This information can be included in the experimental design via the specification of the vector β . By specifying prior marginal utilities different from zero for a particular attribute, the experimenter can express his/her a priori belief about the respondents' preferences regarding the attribute's levels. This prior information might have an important impact on the structure of the design, especially on the number of times each of the attributes is held constant.

In order to study the effect of the prior parameter vector β , we selected a fixed experimental configuration: a SC experiment with 6 three-level attributes, 15 choice situations, 2 alternatives per choice situation and a profile strength of 2. We also defined two factors in order to systematically generate the vectors β to be studied: the number of attributes the prior marginal utilities of which are different from zero (in other words, the number of attributes for which the experimenter has a priori knowledge of the respondents' preferences) and the magnitude of the prior marginal utilities associated to those attributes (in other words, the experimenter's belief about the importance of those attributes and corresponding levels). In our study, we assumed ordinal attributes the levels of which are ordered from least preferred to most preferred, and we chose the non-zero prior marginal utilities to be equally spaced around zero. This assumption, together with the use of effects-type coding, requires specifying the prior marginal utilities of the least preferred level only. Given a marginal utility i for the least preferred level, the marginal utilities for the second and third levels are 0 and $-i$, respectively. In this experiment, we used three levels for the first factor and four levels for the second factor, resulting in $3 \times 4 = 12$ prior parameter vectors β . The levels of each factor are shown in Table 6. Observe that the last level of the marginal utilities (-2) is quite extreme, since it is not commonly observed in real experimental situations involving effects-type coded categorical attributes.

Using the two-stage algorithm, the integrated algorithm and the extensive algorithm, locally optimal designs were generated for each of the 12 non-zero parameter vectors β defined by the factor levels in Table 6 and for $\beta = \mathbf{0}_{12}$. Table 7 shows the \mathcal{D}_P -efficiencies of the designs obtained using each algorithm, together with the number of times each attribute is held constant.

The two-stage algorithm generates designs in which all the attributes are held constant an equal number of times. This is because the prior parameter vector β is not taken into account during the first stage of the algorithm, where the constant attributes are defined for each choice situation. In the designs generated by the integrated and the extensive algorithm, the attributes with non-zero prior marginal utilities are held constant more frequently. The larger the magnitude of the marginal utilities, the larger the number of times these attributes are held constant. This phenomenon is more pronounced when there are fewer attributes with non-zero prior marginal utilities. The lack of balance in the number of times attributes are held constant can be explained as follows: when the marginal utilities of an attribute are large, the choice probabilities

Table 8

Factors and levels tested in Section 6.4 to study the influence of the model parameters' prior variance on the performance of the two-stage, the integrated and the extensive algorithm.

Parameter	Levels
Number of attributes with non-zero variance	2, 4, 6
Variance	0.1, 0.2, 0.3, 0.4

Table 9

Performance comparison of the two-stage, the integrated and the extensive algorithm using 13 prior variance specifications for the parameters.

Variances	Two-stage		Integrated		Extensive	
	\mathcal{D}_B -eff.	Num. const.	\mathcal{D}_B -eff.	Num. const.	\mathcal{D}_B -eff. ^a	Num. const.
0.0 0.0 0.0 0.0 0.0 0.0	99.17	5 5 5 5 5 5	99.25	5 5 5 6 4 5	100.00	5 5 6 6 4 4
0.1 0.1 0.0 0.0 0.0 0.0	98.43	5 5 5 5 5 5	99.75	6 7 6 4 3 4	100.00	6 6 4 5 5 4
0.1 0.1 0.1 0.1 0.0 0.0	98.84	5 5 5 5 5 5	99.75	6 5 6 6 4 3	100.00	4 5 7 7 4 3
0.1 0.1 0.1 0.1 0.1 0.1	98.84	5 5 5 5 5 5	99.75	6 4 6 4 5 5	100.00	5 5 5 5 5 5
0.2 0.2 0.0 0.0 0.0 0.0	97.29	5 5 5 5 5 5	99.67	7 7 3 5 4 4	100.00	8 7 4 3 4 4
0.2 0.2 0.2 0.2 0.0 0.0	97.04	5 5 5 5 5 5	99.50	6 6 6 6 2 4	100.00	6 7 6 6 2 3
0.2 0.2 0.2 0.2 0.2 0.2	98.84	5 5 5 5 5 5	99.34	6 4 5 5 5 5	100.00	6 4 5 5 5 5
0.3 0.3 0.0 0.0 0.0 0.0	94.96	5 5 5 5 5 5	99.34	8 9 4 2 3 4	100.00	9 9 3 2 3 4
0.3 0.3 0.3 0.3 0.0 0.0	95.92	5 5 5 5 5 5	99.42	7 7 6 7 1 2	100.00	7 7 7 6 0 3
0.3 0.3 0.3 0.3 0.3 0.3	99.42	5 5 5 5 5 5	99.34	5 5 5 5 5 5	100.00	5 5 4 6 5 5
0.4 0.4 0.0 0.0 0.0 0.0	93.16	5 5 5 5 5 5	99.83	9 9 4 2 3 3	100.00	9 9 2 4 3 3
0.4 0.4 0.4 0.4 0.0 0.0	93.94	5 5 5 5 5 5	99.42	7 7 7 7 1 1	100.00	8 7 7 7 0 1
0.4 0.4 0.4 0.4 0.4 0.4	99.09	5 5 5 5 5 5	99.00	5 5 5 4 6 5	100.00	5 5 6 3 5 6
Average	97.30		99.49		100.00	

^a Since the designs generated by the extensive algorithm were considered to be optimal, these efficiencies are set to 100%.

for alternatives that differ with respect to that attribute are extreme. Therefore, choice situations for which that attribute is not held constant are not very informative. As a result, such choice situations are not selected by the \mathcal{D} -optimality criterion.

The lack of balance in the number of times the attributes are held constant in the designs produced by the integrated and the extensive algorithm is inversely related to the performance of the two-stage algorithm. The larger the lack of balance in the pattern of constant attributes in the designs produced by the integrated algorithms, the poorer the quality of the designs produced by the two-stage algorithm. For example, when there are only two attributes with non-zero prior marginal utilities and the marginal utility of the least preferred level is equal to -1.5 , the design generated by the two-stage algorithm is 8.91% less \mathcal{D}_B -efficient than that generated by the integrated algorithm and 8.99% less \mathcal{D}_B -efficient than that generated by the extensive algorithm. This illustrates the benefits of considering the prior parameter vector β during the optimization of the set of constant attributes while generating locally optimal designs.

6.4. Influence of the model parameters' prior variance

Bayesian optimal designs are based on a prior distribution of parameter values with mean vector β_0 and variance-covariance matrix Σ_0 . The larger the variance specified for a parameter, the larger the experimenter's degree of uncertainty about its value. In order to study the influence of the prior variance, we conducted an experiment with characteristics similar to the one described in Section 6.3. We used the same design configuration, and a fixed vector of prior means, $\beta_0 = \mathbf{0}_{12}$. We also defined two factors in order to generate the variance-covariance matrices Σ_0 to be studied: the number of attributes with non-zero variances for the parameters (in other words, the number of attributes for which the experimenter is uncertain about the true marginal utilities) and the magnitude of the variances of the parameters. We considered three levels for the first factor and four levels for the second factor, and hence, $3 \times 4 = 12$ different configurations of the variance matrix Σ_0 . The levels for each factor are shown in Table 8. In order to obtain equal variances for all the parameters, we followed the recommendations provided by Kessels et al. (2008). For each attribute f , we specified negative covariances between its $L_f - 1$ marginal utilities corresponding to a correlation coefficient of $-1/(L_f - 1)$.

The two-stage algorithm, the integrated algorithm and the extensive algorithm were executed in order to generate Bayesian optimal designs using each of the 12 variance-covariance matrices Σ_0 generated. An additional control instance was included in the study for which all parameters had a variance equal to zero. This instance assumes complete certainty about the parameter values, and the design generated is equivalent to the UN design. Table 9 shows the \mathcal{D}_B -efficiencies of the designs obtained by each algorithm.

The integrated and the extensive algorithm generate designs in which the attributes with a large prior variance are held constant more often. The larger the prior variance of the parameters, the larger the number of times the attributes are held

Table 10
Attributes studied in the SC experiment of Bliemer and Rose (2011).

Identifier	Attribute	Number of levels	Number of parameters
1	Airline	6	5
2	Ticket price	3	1
3	Departure time	3	2
4	Transfer time	3	1
5	Egress price	2 × 3	1
6	Egress time	2 × 3	1

constant. This difference is more prominent when the number of attributes with non-zero prior parameter variances is small. These results are in line with those discussed by Kessels et al. (2011a,c). The imbalance is a result of the sampling procedure from the normal distribution of parameter values. A prior distribution with large variances for some parameters covers some scenarios where there are strong preferences for certain attribute levels. The larger the variance of the distribution, the larger some of the parameter values sampled by the radial-spherical integration approach and the larger the corresponding marginal utilities. As explained in Section 6.3, attributes with large prior marginal utilities lead to many choice situations for which these attributes are held constant.

As previously stated, the prior variance is a measure of uncertainty about the true values of the parameters associated with an attribute. For that reason, the fact that attributes with large prior variances are held constant more often comes as a surprise. As a matter of fact, the larger the uncertainty about an attribute, the more information the experimenter hopes to collect about that attribute and thus the lower the number of times the attribute may be expected to be held constant in the design. Nevertheless, our computational experiments show the opposite.

7. SC experiments with partial profiles in transportation

In this section, we illustrate the usefulness of the integrated algorithm for SC experiments in transportation by means of two examples.

7.1. Designs with partial profiles

We revisit the SC experiment carried out by Bliemer and Rose (2011) in order to show the benefits of considering partial profiles in a real-life scenario. Bliemer and Rose (2011) studied the air travel choices that people make when flying to large cities. In these cases, it is often possible for travellers to either fly directly to the main airport closest to the city, or fly to a regional airport located farther away from the city for a lower fare. The main goal of the experiment was therefore to examine the trade-offs that travellers make between ticket price and overall travel time.

Bliemer and Rose (2011) framed their study in the context of a holiday trip from Amsterdam to Barcelona: the origin airport was Schipol Airport (located in Amsterdam), while the destination airport was either the main Barcelona Airport or the regional Girona Airport (located around 100 km away from Barcelona). In the experiment, respondents were presented with an online survey involving the choice of a flight ticket among five possible alternatives. Respondents were asked to provide their preference for six choice situations, along with their personal information (gender, income, etc.). The ticket's characteristics or attributes in the SC experiment, along with their numbers of levels and the numbers of parameters they involve, are shown in Table 10. Note that attributes 2, 4, 5 and 6 only involve a single parameter since they were modelled as continuous variables, while the other attributes were assumed to be categorical. Note also that there is an implicit factor, the *destination airport* (either Barcelona Airport or Girona Airport), that is not shown in Table 10 but that was taken into account by Bliemer and Rose (2011). Depending on the destination airport, different values for attributes 5 and 6 were selected in the experimental design. Since Girona Airport is located farther away from the city, the values of egress price and egress time were considered to be larger for tickets to this destination (compared to those to Barcelona Airport). For example, when the destination airport was Barcelona, the possible values of the egress price (attribute 5) were 1, 3 and 5 euros; conversely, when the destination airport was Girona, the possible values of the egress price were 9, 12 and 15 euros. In practice, this additional constraint caused each destination airport to have a different set of three possible values for attributes 5 and 6, respectively.

The SC experiment was carried out in three separate stages, each of which implemented a different experimental design. The main objective of the authors was to study the influence of the design on the resulting parameter estimates. The experimental designs investigated were the following: an orthogonal design involving 108 choice situations grouped in 18 blocks of size 6, a Bayesian optimal design also involving 108 choice situations and 18 blocks of size 6, and a Bayesian optimal design involving only 18 choice situations and 3 blocks of size 6. The vector of prior means and the prior variance-covariance matrix used for the generation of the Bayesian optimal designs were obtained via a smaller pilot study.

Bliemer and Rose (2011) report that the Bayesian optimal designs led to parameter estimates with lower standard errors than those produced by the orthogonal design. Moreover, the authors also observed that different designs yielded somewhat different estimates. They believe this to be due to the fact that dominant alternatives in the design might bias the estimation

Table 11

Number of times each attribute is held constant in the designs generated by the integrated algorithm for the SC experiment of [Bliemer and Rose \(2011\)](#).

(a) 18 choice situations							(b) 108 choice situations						
Design	Attribute						Design	Attribute					
	1	2	3	4	5	6		1	2	3	4	5	6
Utility-neutral	0	9	0	9	9	9	Utility-neutral	0	54	0	54	54	54
Locally optimal	0	8	2	9	9	8	Locally optimal	0	52	8	52	52	52
Bayesian	0	9	2	9	8	8	Bayesian	0	51	9	53	52	51

of the parameters. In this sense, Bayesian optimal designs are preferable since they contain fewer choice situations with this kind of alternatives. The Bayesian optimal designs considered by [Bliemer and Rose \(2011\)](#), however, are still vulnerable to bias due to the presence of dominant attributes. In fact, the prior estimates obtained during the pilot study revealed that attribute 3 (departure time) had a crucial influence on the traveller's choice behaviour, while attributes 2, 4 and 6 (ticket price, transfer time and egress time) also had a significant impact. By holding some of these attributes constant in the choice situations of the design, it would have been possible to mitigate any bias due to potential dominant attributes.

We made use of the integrated algorithm in order to investigate the structure of optimal designs with partial profiles for the SC experiment of [Bliemer and Rose \(2011\)](#). For the sake of clarity, we obviate the destination airport and the restriction that this factor imposes on the set of possible values for attributes 5 and 6. We assume, instead, that both attributes have three generic levels representing low, medium and high values, and that these levels are independent from the destination. This small modification causes the experimental scenario to be less restrictive than the original one, while keeping it useful for illustrative purposes. For this scenario, we generated two sets of designs: one set with designs involving 18 choice situations and another set with designs involving 108 choice situations. We generated UN optimal designs, and locally and Bayesian optimal designs using the same prior values as those considered in the original paper. Unlike the designs generated by [Bliemer and Rose \(2011\)](#), our designs do not account for any blocking configuration and do not impose any attribute level balance.

[Table 11](#) shows the number of times each attribute is held constant in each of the designs we generated. The set of designs involving 18 choice situations is shown in [Table 12](#). Observe that in the choice situations of the UN designs, attributes 1 and 3 are never held constant. This is mainly because these attributes involve a larger number of parameters than the other attributes (see [Table 10](#)). Therefore, they need to be varied more often in order to allow for good estimates of their parameter values. In the locally optimal and Bayesian optimal designs, attribute 3 is held constant in a few choice situations since this attribute was identified to be very influential during the pilot study. These choice situations allow to collect valuable information because they force the respondents to focus on the other attributes, and pay less attention to the (potentially) dominant one.

7.2. The integrated algorithm

We revisit the SC experiment carried out by [Kupfer et al. \(2016\)](#) in order to show the benefits of using the integrated algorithm in a real-life scenario. [Kupfer et al. \(2016\)](#) study the airport competition for air cargo in Europe. They identify the most important characteristics that airlines take into account when selecting an airport for their freighter operations, and thereby offer better insights to airports and policy makers about how to attract new service providers. [Table 13](#) shows the airport characteristics considered as categorical attributes in the SC experiment, along with their numbers of levels. The preferences of the airlines were collected by means of questionnaires consisting of 20 choice situations, each involving two alternative airports as profiles. In order to limit the cognitive burden imposed on the respondents, each choice situation showed only four of the six attributes studied. In other words, each choice situation was composed of partial profiles in which two attributes were held constant.

[Kupfer et al. \(2016\)](#) expected the origin-destination demand (attribute 6) to dominate the airlines' choices. For that reason, their SC experiment involved choice situations in which attribute 6 was held constant more often. To achieve that, the authors generated two different Bayesian \mathcal{D} -optimal designs using the two-stage algorithm of [Kessels et al. \(2015\)](#): a 6-attribute design involving 40 choice situations and two constant attributes, and a 5-attribute design (excluding attribute 6) involving 40 choice situations and one constant attribute. Both designs consist of four different blocks of 10 choice situations, where one block from the 6-attribute design and one block from the 5-attribute design were assigned to each respondent. In doing so, the authors ensured that attribute 6 was held constant in at least 50% of the choice situations that composed the resulting four questionnaires.

The design for the SC experiment of [Kupfer et al. \(2016\)](#) can be improved by using the integrated algorithm. First, as shown in [Sections 6.3](#) and [6.4](#), the integrated algorithm generates Bayesian \mathcal{D} -optimal designs that are more efficient than those generated by the two-stage algorithm. In fact, the 5-attribute and the 6-attribute designs generated by the two-stage algorithm (considering the same prior distribution of parameter values) are between 1% and 2% less \mathcal{D}_B -efficient than the designs generated by the integrated algorithm. This is mainly due to the ability of the integrated algorithm to generate

Table 12
 Designs for the SC experiment of [Bliemer and Rose \(2011\)](#) involving 18 choice situations, 5 alternatives per choice situation, 1 six-level attribute and 5 three-level attributes.

CS	Utility-neutral					Locally optimal					Bayesian							
	1	2	3	4	5	6	1	2	3	4	5	6	1	2	3	4	5	6
1	3	*	1	*	1	3	2	*	*	3	3	1	1	*	*	1	3	1
	2	*	2	*	3	1	3	*	*	1	3	3	4	*	*	1	1	3
	5	*	2	*	1	3	4	*	*	3	1	3	6	*	*	3	3	1
	1	*	1	*	3	3	6	*	*	3	3	3	2	*	*	3	1	3
	6	*	3	*	1	1	1	*	*	1	1	1	3	*	*	1	1	3
2	4	*	3	*	1	3	1	*	2	*	1	3	2	*	1	*	3	3
	3	*	1	*	3	1	6	*	1	*	3	1	5	*	1	*	3	1
	2	*	3	*	3	3	5	*	3	*	1	1	6	*	3	*	1	3
	1	*	2	*	1	1	3	*	3	*	3	1	1	*	3	*	3	3
	6	*	1	*	3	3	4	*	1	*	3	3	3	*	2	*	1	1
3	3	*	3	1	*	3	6	*	1	*	3	3	5	*	2	*	1	3
	1	*	3	3	*	3	1	*	3	*	3	3	4	*	3	*	3	1
	5	*	2	3	*	1	5	*	3	*	1	3	2	*	3	*	3	1
	2	*	1	1	*	3	2	*	3	*	3	1	1	*	1	*	3	3
	4	*	2	1	*	1	4	*	2	*	1	1	6	*	1	*	1	1
4	6	*	1	1	*	3	4	*	3	*	3	1	2	*	3	*	1	1
	3	*	3	3	*	1	2	*	1	*	3	3	4	*	2	*	3	1
	4	*	3	3	*	3	5	*	1	*	3	3	5	*	1	*	1	3
	1	*	2	1	*	1	1	*	3	*	1	3	1	*	3	*	3	3
	2	*	2	3	*	3	6	*	2	*	1	1	3	*	1	*	1	3
5	6	*	2	3	*	3	2	*	2	*	1	3	4	*	1	1	*	1
	2	*	3	1	*	3	1	*	1	*	3	1	3	*	1	1	*	3
	5	*	3	3	*	3	3	*	3	*	3	1	2	*	3	3	*	1
	1	*	1	3	*	1	4	*	1	*	3	1	6	*	2	3	*	3
	4	*	2	1	*	1	5	*	3	*	1	3	5	*	3	1	*	1
6	4	*	2	3	1	*	6	*	2	3	*	3	5	*	2	3	*	1
	5	*	3	1	3	*	4	*	1	1	*	1	1	*	1	1	*	3
	2	*	1	3	1	*	5	*	3	1	*	3	4	*	1	3	*	3
	1	*	1	3	3	*	1	*	3	3	*	1	3	*	3	1	*	3
	3	*	3	1	3	*	3	*	1	1	*	1	6	*	3	1	*	1
7	5	*	1	1	1	*	2	*	3	1	*	1	1	*	1	1	3	*
	6	*	2	3	3	*	6	*	1	1	*	1	5	*	1	3	3	*
	3	*	3	3	3	*	3	*	3	1	*	3	3	*	3	3	3	*
	1	*	3	3	1	*	4	*	2	3	*	3	4	*	3	3	1	*
	4	*	1	1	1	*	5	*	1	3	*	1	2	*	2	1	1	*
8	5	*	1	3	3	*	2	*	3	1	*	3	6	*	2	1	3	*
	3	*	3	3	3	*	5	*	2	3	*	1	2	*	3	3	1	*
	1	*	2	1	1	*	6	*	3	3	*	1	5	*	3	1	3	*
	6	*	1	1	1	*	3	*	1	1	*	3	4	*	3	3	1	*
	4	*	2	3	3	*	4	*	1	1	*	3	1	*	1	3	1	*
9	4	*	2	1	3	*	2	1	*	3	*	1	6	*	3	3	1	*
	2	*	3	1	3	*	3	1	*	3	*	3	3	*	1	3	3	*
	5	*	2	3	1	*	1	3	*	1	*	3	5	*	3	3	3	*
	1	*	1	3	1	*	4	3	*	1	*	1	4	*	3	1	3	*
	6	*	1	3	1	*	5	3	*	3	*	3	1	*	2	1	1	*
10	4	3	1	*	*	3	1	3	2	*	*	1	5	3	*	*	3	3
	2	1	1	*	*	1	4	3	1	*	*	3	3	1	*	*	3	3
	5	3	3	*	*	1	6	1	1	*	*	3	4	1	*	*	1	1
	3	1	2	*	*	3	5	1	3	*	*	3	1	3	*	*	1	1
	6	3	2	*	*	3	3	1	3	*	*	1	2	3	*	*	3	3
11	5	3	2	*	*	3	6	3	3	*	1	*	4	1	1	*	*	3
	2	3	1	*	*	1	5	1	1	*	1	*	3	3	2	*	*	1
	4	1	3	*	*	1	3	3	2	*	3	*	6	1	3	*	*	3
	6	1	3	*	*	1	1	1	3	*	3	*	5	1	3	*	*	3
	1	1	2	*	*	3	2	1	1	*	1	*	2	3	1	*	*	1
12	5	1	1	*	*	3	5	3	1	*	1	*	1	3	3	*	*	3
	6	3	3	*	*	1	6	3	3	*	3	*	2	1	2	*	*	3
	4	3	1	*	*	1	2	1	3	*	1	*	5	1	1	*	*	1
	2	3	2	*	*	3	1	1	2	*	3	*	3	3	1	*	*	1
	3	1	2	*	*	1	3	3	1	*	1	*	6	3	3	*	*	1

(continued on next page)

Table 12 (continued)

CS	Utility-neutral						Locally optimal						Bayesian					
	1	2	3	4	5	6	1	2	3	4	5	6	1	2	3	4	5	6
13	1	1	1	*	*	3	2	3	2	*	3	*	4	3	2	*	*	3
	3	3	2	*	*	3	6	1	3	*	1	*	3	1	3	*	*	1
	2	1	2	*	*	1	4	1	3	*	3	*	1	1	3	*	*	1
	4	1	3	*	*	3	1	3	3	*	1	*	2	1	1	*	*	1
	5	3	1	*	*	1	3	1	1	*	1	*	5	3	3	*	*	3
14	6	3	2	*	3	*	2	3	1	*	1	*	1	1	3	*	1	*
	5	3	3	*	3	*	6	3	3	*	1	*	2	3	2	*	3	*
	4	1	1	*	3	*	3	1	1	*	1	*	5	1	3	*	1	*
	2	3	3	*	1	*	4	3	3	*	3	*	4	3	1	*	1	*
	3	1	2	*	1	*	5	1	2	*	3	*	3	1	1	*	3	*
15	2	1	1	*	3	*	6	3	1	3	*	*	1	3	3	*	3	*
	1	3	2	*	3	*	4	3	3	1	*	*	4	1	2	*	3	*
	6	1	3	*	1	*	2	1	2	1	*	*	6	3	1	*	1	*
	3	3	1	*	1	*	1	3	1	3	*	*	3	1	3	*	1	*
	5	1	3	*	1	*	3	1	3	3	*	*	5	3	1	*	1	*
16	6	3	2	*	3	*	2	1	3	3	*	*	1	3	2	3	*	*
	5	1	1	*	3	*	6	3	3	1	*	*	3	1	3	1	*	*
	1	1	3	*	3	*	3	3	2	3	*	*	2	1	1	3	*	*
	2	3	3	*	1	*	1	1	1	1	*	*	6	1	1	1	*	*
	3	1	2	*	1	*	5	1	1	1	*	*	4	3	3	1	*	*
17	4	3	1	3	*	*	1	3	1	3	*	*	5	3	1	1	*	*
	5	1	2	1	*	*	2	3	1	3	*	*	6	3	3	1	*	*
	3	3	1	3	*	*	6	1	2	1	*	*	4	3	1	3	*	*
	6	1	3	3	*	*	4	1	3	3	*	*	1	1	2	3	*	*
	1	3	3	1	*	*	5	3	3	1	*	*	2	1	3	1	*	*
18	1	3	3	1	*	*	4	1	3	3	*	*	6	1	1	3	*	*
	4	3	3	3	*	*	5	3	2	1	*	*	4	3	3	3	*	*
	3	3	1	1	*	*	2	3	3	1	*	*	3	3	3	3	*	*
	6	1	1	1	*	*	1	1	1	3	*	*	5	1	2	1	*	*
	2	1	2	3	*	*	3	1	3	3	*	*	2	3	1	1	*	*

Table 13
Attributes studied in the SC experiment of Kupfer et al. (2016).

Identifier	Attribute	Number of levels
1	Night-time restrictions	3
2	Airport experience with cargo	3
3	Presence of forwarders	3
4	Presence of passenger airlines	3
5	Airport charges (including handling)	5
6	Origin-destination demand	5

Table 14
Number of times each attribute is held constant in the designs generated by the two-stage algorithm and the designs generated by the integrated algorithm for the SC experiment of Kupfer et al. (2016).

	(a) 6-attribute designs						(b) 5-attribute designs						
	Algorithm	Attribute					Algorithm	Attribute					
		1	2	3	4	5	6		1	2	3	4	5
Two-stage		16	16	17	17	7	7	Two-stage	10	10	10	10	0
Integrated		21	20	17	19	0	3	Integrated	11	10	11	8	0

designs with a better choice for the constant attributes. The structural differences between the designs are reported in Table 14, which shows the number of times each attribute is held constant in each design. Observe that the differences are more substantial between the 6-attribute designs.

A second advantage of using the integrated algorithm is that an ad hoc design construction, involving two different subdesigns (one with and one without the sixth attribute), is no longer needed. If an experimenter expects an attribute to be dominant (as was the case in Kupfer et al. (2016)), and would like to have a design in which that particular attribute is held constant more often, it is not necessary to combine two different designs in an ad hoc fashion. The experimenter can

express his/her a priori belief by specifying larger marginal utilities for that particular attribute (reflecting the attribute's dominance) and generating a Bayesian optimal design: the integrated algorithm will then automatically generate a design in which this attribute is held constant more often.

8. Summary and conclusion

In this paper, we have dealt with the generation of \mathcal{D} -optimal designs for SC experiments with partial profiles. The theoretical approaches available in the literature allow the generation of optimal designs for a restricted set of experimental situations. In response to these limitations, [Kessels et al. \(2011b, 2015\)](#) propose more flexible two-stage algorithms. The first stage of the algorithms determines the set of constant attributes for each choice situation, and the second stage determines the levels of the varying attributes. These algorithms, however, show some limitations due to the fact that two optimization steps are carried out independently. In this paper, we propose an integrated algorithm that determines the set of constant attributes and the levels of the varying attributes simultaneously. The algorithm modifies the design, choice situation by choice situation, in order to optimize the value of the optimality criterion. The integrated algorithm, together with a simple multi-start strategy, generates designs with a very good quality in a short execution time. Because the algorithm does not always find the optimal designs, we also propose an extension to the integrated algorithm. This more extensive algorithm has a considerably longer execution time; however, it does find the optimal designs for 98% of the benchmark experiments analytically derived. As a result, the extensive algorithm is very powerful besides being very flexible.

The extensive set of computational experiments carried out shows the benefits of performing a simultaneous optimization. The designs generated by the integrated algorithm and the extensive version consistently outperform those generated by the two-stage algorithm. This is mainly because the first stage of the two-stage algorithm, which defines the structure of constant attributes, does not consider all the characteristics of the experiment. This leads the two-stage algorithm to generate designs with structures that are sometimes very different from those of the true optimal designs. The performance difference is even more prominent when creating locally optimal and Bayesian optimal designs. The \mathcal{D}_B -optimality criterion leads to designs for which the attributes with large prior values for the marginal utilities or large prior variances are held constant more often. The former result should not come as a surprise, since large prior values for the marginal utilities indicate that the experimenter is certain about strong preferences regarding these attributes. Therefore, it is reasonable to hold these attributes constant more often. However, the latter result is counterintuitive. By specifying a large prior variance for an attribute's marginal utilities, the experimenter expresses a large degree of uncertainty about the true values of the attribute's marginal utilities. Therefore, the experimenter certainly hopes that attribute to be held constant less often in order to collect more information about it. These results cast some doubt about the way in which the Bayesian \mathcal{D} -optimality criterion is implemented in the literature on SC experiments. This should encourage the study of alternative criteria in future research.

In this paper, we have focused on the optimal design of SC experiments involving unlabelled alternatives, categorical and continuous attributes that are fixed across alternatives and the estimation of the MNL model involving main effects. The integrated algorithm, however, is not limited to this experimental scenario only. It can be adapted without major difficulties to handle other discrete choice models and account for special characteristics, such as labelled alternatives. Nevertheless, there are two major considerations that need to be taken into account, one of which related to the use of partial profiles and the other one related to the integrated algorithm itself.

- The use of partial profiles implicitly assumes that an attribute can be held constant because its levels are comparable across all the alternatives in a choice situation. In other words, it is assumed that the sets of possible attribute levels are the same for every alternative, and that the influence of the attribute on the utility is quantified by means of a single set of parameters. If an attribute does not satisfy these properties (it is not present in all the alternatives of the choice situation, the sets of possible levels are not the same for all the alternatives, or the impact of its attribute levels on the utility is modelled by means of different sets of parameters), it should not be held constant in a choice situation. The integrated algorithm should be modified in order to always consider this kind of attribute as a varying attribute.
- Even though the integrated algorithm can be extended to consider more complex discrete choice models, its optimization process is tailored to handle models with main effects only. Under the main-effects model, the levels of the attributes that are held constant in the choice situations do not impact the respondents' choices or the quality of the design. For this reason, in many experimental scenarios, the levels of the constant attributes are not shown. Respondents are told in advance to only consider the attribute levels that are displayed. As a consequence, when generating an optimal design, the integrated algorithm can simply fix the constant attributes to their first level (see [Algorithm 3](#)). In contrast, under a model involving interaction effects, the levels of the constant attributes do have an impact on the respondents' choices and the quality of the design. Therefore, all attribute levels, including those that are constant in the choice situations, have to be shown to the respondents. Not doing so can lead to biased utility estimates. This is mainly because the respondents' choices might differ depending on the assumptions they make about the levels of the missing attributes ([Chrzan, 2010](#)). For these reasons, in order to generate designs for models including interaction terms, it would be necessary to extend the algorithm in order to also optimize the levels of the constant attributes. This modification is relatively straightforward to implement and will be explored in future research.

Supplementary materials

The designs generated for the benchmark instances listed by Großmann et al. (2009) and the SC experiment of Bliemer and Rose (2011) are available as supplementary materials on the website <http://antor.uantwerpen.be/optimal-designs-TRB>.

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